

WHAT IS VOLATILITY?

PROBABILITY IS IRRELEVANT IN OPTION-PRICING

Tommi Sottinen

University of Vaasa

2nd October 2008

ABSTRACT

Volatility plays a fundamental rôle in econometric modelling and in option pricing. However, it seems that it is not clear what it is.

ABSTRACT

Volatility plays a fundamental rôle in econometric modelling and in option pricing. However, it seems that it is not clear what it is.

To illustrate the problem we construct a toy-model that incorporates **LONG-RANGE DEPENDENCE** and **HEAVY TAILS** to the standard Black–Scholes model while keeping the replication prices of options unchanged.

ABSTRACT

Volatility plays a fundamental rôle in econometric modelling and in option pricing. However, it seems that it is not clear what it is.

To illustrate the problem we construct a toy-model that incorporates **LONG-RANGE DEPENDENCE** and **HEAVY TAILS** to the standard Black–Scholes model while keeping the replication prices of options unchanged.

So, the volatility as the pricing parameter is the same as in the classical Black–Scholes model, but the historical volatility (standard deviation) is not the same as in the Black–Scholes model. Indeed, the historical volatility may not even exist.

ABSTRACT

The moral of the story is

- The historical volatility and the implied volatility need not have anything in common.
- The probabilistic properties of the pricing model are mostly irrelevant in option-pricing.

ABSTRACT

The moral of the story is

- The historical volatility and the implied volatility need not have anything in common.
- The probabilistic properties of the pricing model are mostly irrelevant in option-pricing.

The talk is based on C. BENDER, T. SOTTINEN, and E. VALKEILA (2008): Pricing by hedging and no-arbitrage beyond semimartingales. *Finance and Stochastics*, forthcoming.

OUTLINE

- 1 TOY MODEL
- 2 HEDGING WITH QUADRATIC VARIATION
- 3 TOY MODEL'S VOLATILITY
- 4 REFERENCES

OUTLINE

1 TOY MODEL

2 HEDGING WITH QUADRATIC VARIATION

3 TOY MODEL'S VOLATILITY

4 REFERENCES

TOY MODEL

INGREDIENTS

We consider **DISCOUNTED MARKETS** with one risky asset given by the mixed model

$$S_t = s_0 \exp \left\{ \mu t + \sigma W_t - \frac{\sigma^2}{2} t + \delta B_t^H - I_t^{\alpha_1} + I_t^{\alpha_2} \right\},$$

where

TOY MODEL

INGREDIENTS

We consider **DISCOUNTED MARKETS** with one risky asset given by the mixed model

$$S_t = s_0 \exp \left\{ \mu t + \sigma W_t - \frac{\sigma}{2} t + \delta B_t^H - I_t^{\alpha_1} + I_t^{\alpha_2} \right\},$$

where

- B^H is a **FRACTIONAL BROWNIAN MOTION** with Hurst index $H > 0.5$.

TOY MODEL

INGREDIENTS

We consider **DISCOUNTED MARKETS** with one risky asset given by the mixed model

$$S_t = s_0 \exp \left\{ \mu t + \sigma W_t - \frac{\sigma^2}{2} t + \delta B_t^H - I_t^{\alpha_1} + I_t^{\alpha_2} \right\},$$

where

- B^H is a **FRACTIONAL BROWNIAN MOTION** with Hurst index $H > 0.5$.
- $I_t^{\alpha_i}$'s are **INTEGRATED COMPOUND POISSON PROCESSES** with positive heavy-tailed jumps:

$$I_t^{\alpha_i} = \int_0^t \sum_{k: \tau_k^i \leq s} U_k^i ds,$$

τ_k^i 's are Poisson arrivals and $\mathbf{P}[U_k^i > x] \sim x^{-\alpha_i}$.

TOY MODEL

INGREDIENTS

We consider **DISCOUNTED MARKETS** with one risky asset given by the mixed model

$$S_t = s_0 \exp \left\{ \mu t + \sigma W_t - \frac{\sigma}{2} t + \delta B_t^H - I_t^{\alpha_1} + I_t^{\alpha_2} \right\},$$

where

- B^H is a **FRACTIONAL BROWNIAN MOTION** with Hurst index $H > 0.5$.
- I^{α_i} 's are **INTEGRATED COMPOUND POISSON PROCESSES** with positive heavy-tailed jumps:

$$I_t^{\alpha_i} = \int_0^t \sum_{k: \tau_k^i \leq s} U_k^i ds,$$

τ_k^i 's are Poisson arrivals and $\mathbf{P}[U_k^i > x] \sim x^{-\alpha_i}$.

- W , B^H , I^{α_1} , and I^{α_2} are independent.

TOY MODEL

STYLIZED FACTS

Stylized facts for the returns R_t in the mixed model:

Stylized facts for the returns R_t in the mixed model:

1 LONG-RANGE DEPENDENCE: If I^{α_i} 's are in L^2 then

$$\mathbf{Cor}[R_1, R_t] \sim \delta^2 H(2H - 1)t^{2H-2}.$$

Stylized facts for the returns R_t in the mixed model:

1 LONG-RANGE DEPENDENCE: If I^{α_i} 's are in L^2 then

$$\mathbf{Cor}[R_1, R_t] \sim \delta^2 H(2H - 1)t^{2H-2}.$$

2 HEAVY TAILS: $\mathbf{P}[-R_t > x] \gtrsim x^{-\alpha_1}$ and $\mathbf{P}[R_t > x] \gtrsim x^{-\alpha_2}$.

TOY MODEL

STYLIZED FACTS

Stylized facts for the returns R_t in the mixed model:

1 LONG-RANGE DEPENDENCE: If I^{α_i} 's are in L^2 then

$$\mathbf{Cor}[R_1, R_t] \sim \delta^2 H(2H - 1)t^{2H-2}.$$

2 HEAVY TAILS: $\mathbf{P}[-R_t > x] \gtrsim x^{-\alpha_1}$ and $\mathbf{P}[R_t > x] \gtrsim x^{-\alpha_2}$.

3 GAIN/LOSS ASYMMETRY: Obvious if $\alpha_1 < \alpha_2$.

TOY MODEL

STYLIZED FACTS

Stylized facts for the returns R_t in the mixed model:

1 LONG-RANGE DEPENDENCE: If I^{α_i} 's are in L^2 then

$$\mathbf{Cor}[R_1, R_t] \sim \delta^2 H(2H - 1)t^{2H-2}.$$

2 HEAVY TAILS: $\mathbf{P}[-R_t > x] \gtrsim x^{-\alpha_1}$ and $\mathbf{P}[R_t > x] \gtrsim x^{-\alpha_2}$.

3 GAIN/LOSS ASYMMETRY: Obvious if $\alpha_1 < \alpha_2$.

4 JUMPS: No, but can you tell the difference between jumps and heavy tails from a discrete data?

Stylized facts for the returns R_t in the mixed model:

1 LONG-RANGE DEPENDENCE: If I^{α_i} 's are in L^2 then

$$\mathbf{Cor}[R_1, R_t] \sim \delta^2 H(2H - 1)t^{2H-2}.$$

2 HEAVY TAILS: $\mathbf{P}[-R_t > x] \gtrsim x^{-\alpha_1}$ and $\mathbf{P}[R_t > x] \gtrsim x^{-\alpha_2}$.

3 GAIN/LOSS ASYMMETRY: Obvious if $\alpha_1 < \alpha_2$.

4 JUMPS: No, but can you tell the difference between jumps and heavy tails from a discrete data?

5 VOLATILITY CLUSTERING: What is volatility? If volatility is **STANDARD DEVIATION**, we can get any kind of volatility structure: Change the Poisson arrivals to clustered arrivals. If volatility (squared) is the so-called **QUADRATIC VARIATION** then it is fixed to constant σ^2 .

OUTLINE

1 TOY MODEL

2 HEDGING WITH QUADRATIC VARIATION

3 TOY MODEL'S VOLATILITY

4 REFERENCES

HEDGING WITH QUADRATIC VARIATION

FORWARD INTEGRALS

The forward integral is economically meaningful in the context of self-financing strategies:

Let (π_n) is a fixed sequence of, say, dyadic partitions of $[0, T]$.

Then the **FORWARD INTEGRAL**

$$\int_0^t \Phi_u dS_u$$

(along the sequence of partitions (π_n)) is the **P**-a.s. forward-sum limit

$$\lim_{n \rightarrow \infty} \sum_{\substack{t_k \in \pi_n \\ t_k \leq t}} \Phi_{t_{k-1}} (S_{t_k} - S_{t_{k-1}})$$

(when it exists).

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Let (π_n) is a fixed sequence of, say, dyadic partitions of $[0, T]$.
Then the **QUADRATIC VARIATION**

$$\langle S \rangle_t$$

(along the sequence of partitions (π_n)) is the **P**-a.s. limit

$$\lim_{n \rightarrow \infty} \sum_{\substack{t_k \in \pi_n \\ t_k \leq t}} (S_{t_k} - S_{t_{k-1}})^2$$

(when it exists).

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Some formulas for Quadratic Variation:

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Some formulas for Quadratic Variation:

1 If $\langle Y \rangle = 0$, then $\langle X + Y \rangle_t = \langle X \rangle_t$;

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Some formulas for Quadratic Variation:

- 1 If $\langle Y \rangle = 0$, then $\langle X + Y \rangle_t = \langle X \rangle_t$;
- 2 If X is differentiable, then $\langle X \rangle_t = 0$;

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Some formulas for Quadratic Variation:

- 1 If $\langle Y \rangle = 0$, then $\langle X + Y \rangle_t = \langle X \rangle_t$;
- 2 If X is differentiable, then $\langle X \rangle_t = 0$;

3

$$\left\langle \int_0^\cdot f(X_u) dX_u \right\rangle_t = \int_0^t f(X_u)^2 d\langle X \rangle_u;$$

HEDGING WITH QUADRATIC VARIATION

QUADRATIC VARIATION

Some formulas for Quadratic Variation:

1 If $\langle Y \rangle = 0$, then $\langle X + Y \rangle_t = \langle X \rangle_t$;

2 If X is differentiable, then $\langle X \rangle_t = 0$;

3

$$\left\langle \int_0^\cdot f(X_u) dX_u \right\rangle_t = \int_0^t f(X_u)^2 d\langle X \rangle_u;$$

4

$$\langle g \circ X \rangle_t = \int_0^t g'(X_u) d\langle X \rangle_u.$$

HEDGING WITH QUADRATIC VARIATION

ITÔ'S LEMMA

THEOREM

Let $f \in \mathcal{C}^{1,2}([0, T], \mathbb{R}_+)$. If S has quadratic variation then we have the Itô formula

$$df(t, S_t) = f_t(t, S)dt + f_x(t, S)dS_t + \frac{1}{2}f_{xx}(t, S_t)d\langle S \rangle_t$$

HEDGING WITH QUADRATIC VARIATION

ITÔ'S LEMMA

THEOREM

Let $f \in \mathcal{C}^{1,2}([0, T], \mathbb{R}_+)$. If S has quadratic variation then we have the Itô formula

$$df(t, S_t) = f_t(t, S)dt + f_x(t, S)dS_t + \frac{1}{2}f_{xx}(t, S_t)d\langle S \rangle_t$$

PROOF.

Taylor is all you need. □

HEDGING WITH QUADRATIC VARIATION

ITÔ'S LEMMA

THEOREM

Let $f \in \mathcal{C}^{1,2}([0, T], \mathbb{R}_+)$. If S has quadratic variation then we have the Itô formula

$$df(t, S_t) = f_t(t, S)dt + f_x(t, S)dS_t + \frac{1}{2}f_{xx}(t, S_t)d\langle S \rangle_t$$

PROOF.

Taylor is all you need. □

REMARK

Itô's formula implies that the forward integral on the right hand side exists and has a continuous modification.

HEDGING WITH QUADRATIC VARIATION

BLACK-SCHOLES BPDE

THEOREM

Let $F(S_T)$ be a European option with maturity T . Let $f(t, S_t)$ satisfy the **BLACK-SCHOLES BPDE**

$$f_t(t, x) + \frac{\sigma^2 x^2}{2} f_{xx}(t, x) = 0, \quad f(T, x) = F(x).$$

Then $f_x(t, S_t)$ is the Delta-hedge for $F(S_T)$ and $f(0, s_0)$ is the price of the option.

HEDGING WITH QUADRATIC VARIATION

BLACK-SCHOLES BPDE

THEOREM

Let $F(S_T)$ be a European option with maturity T . Let $f(t, S_t)$ satisfy the **BLACK-SCHOLES BPDE**

$$f_t(t, x) + \frac{\sigma^2 x^2}{2} f_{xx}(t, x) = 0, \quad f(T, x) = F(x).$$

Then $f_x(t, S_t)$ is the Delta-hedge for $F(S_T)$ and $f(0, s_0)$ is the price of the option.

PROOF.

Note that $d\langle S \rangle_t = \sigma^2 S_t^2 dt$, and then Itô is all you need. □

HEDGING WITH QUADRATIC VARIATION

ON GIRSANOV AND FEYNMAN-KAC

- 1 We did not deal with any Equivalent Martingale Measures here. So, there are no **GIRSANOV RESTRICTIONS** to the drift of S .

HEDGING WITH QUADRATIC VARIATION

ON GIRSANOV AND FEYNMAN-KAC

- 1 We did not deal with any Equivalent Martingale Measures here. So, there are no **GIRSANOV RESTRICTIONS** to the drift of S .
- 2 The **FEYNMAN-KAC** connection to BPDEs tells us that

$$f(t, x) = \mathbf{E} \left[F(\tilde{S}_T) \mid \tilde{S}_t = x \right],$$

where \tilde{S} is the Geometric Brownian Motion. This is true despite of the facts that our toy model is not log-normal, and the returns are not independent.

OUTLINE

- 1 TOY MODEL
- 2 HEDGING WITH QUADRATIC VARIATION
- 3 TOY MODEL'S VOLATILITY**
- 4 REFERENCES

TOY MODEL'S VOLATILITY

HISTORICAL VOLATILITY

Let R_k 's be the log-returns

$$R_k = \log \frac{S_k}{S_{k-1}}, \quad k = 1, 2, \dots$$

■ Then

$$\mathbf{Var}[R_k] = \sigma^2 + \delta^2 + v_{k1}^2 + v_{k2}^2,$$

where $v_{k1}^2, v_{k2}^2 \rightarrow \infty$ (possibly already $+\infty$ for finite k) are the variances of the increments of I^{α_1} and I^{α_2} .

TOY MODEL'S VOLATILITY

HISTORICAL VOLATILITY

Let R_k 's be the log-returns

$$R_k = \log \frac{S_k}{S_{k-1}}, \quad k = 1, 2, \dots$$

- Then

$$\mathbf{Var}[R_k] = \sigma^2 + \delta^2 + v_{k1}^2 + v_{k2}^2,$$

where $v_{k1}^2, v_{k2}^2 \rightarrow \infty$ (possibly already $+\infty$ for finite k) are the variances of the increments of I^{α_1} and I^{α_2} .

- We have that

$$\frac{1}{n} \sum_{k=1}^n \left(R_k - \frac{1}{n} \sum_{k=1}^n R_k \right)^2 \rightarrow \sigma^2 + \delta^2 > \sigma^2$$

if the non-Gaussian parts vanish, and otherwise we do not have convergence at all.

TOY MODEL'S VOLATILITY

HISTORICAL VOLATILITY

Let R_k 's be the log-returns

$$R_k = \log \frac{S_k}{S_{k-1}}, \quad k = 1, 2, \dots$$

- Then

$$\mathbf{Var}[R_k] = \sigma^2 + \delta^2 + v_{k1}^2 + v_{k2}^2,$$

where $v_{k1}^2, v_{k2}^2 \rightarrow \infty$ (possibly already $+\infty$ for finite k) are the variances of the increments of I^{α_1} and I^{α_2} .

- We have that

$$\frac{1}{n} \sum_{k=1}^n \left(R_k - \frac{1}{n} \sum_{k=1}^n R_k \right)^2 \rightarrow \sigma^2 + \delta^2 > \sigma^2$$

if the non-Gaussian parts vanish, and otherwise we do not have convergence at all.

- So, **HISTORICAL VOLATILITY \neq QUADRATIC VARIATION.**

TOY MODEL'S VOLATILITY

IMPLIED VOLATILITY

- The options' prices with toy-model are given by **REPLICATION PRICES**.

TOY MODEL'S VOLATILITY

IMPLIED VOLATILITY

- The options' prices with toy-model are given by **REPLICATION PRICES**.
- So, the **IMPLIED VOLATILITY IS THE QUADRATIC VARIATION σ^2** .

TOY MODEL'S VOLATILITY

IMPLIED VOLATILITY

- The options' prices with toy-model are given by **REPLICATION PRICES**.
- So, the **IMPLIED VOLATILITY IS THE QUADRATIC VARIATION σ^2** .
- The **IMPLIED VOLATILITY IS INDEPENDENT OF THE "SMOOTH" PARTS B^H , I^{α_1} , AND I^{α_2}** .

TOY MODEL'S VOLATILITY

IMPLIED VOLATILITY

- The options' prices with toy-model are given by **REPLICATION PRICES**.
- So, the **IMPLIED VOLATILITY IS THE QUADRATIC VARIATION σ^2** .
- The **IMPLIED VOLATILITY IS INDEPENDENT OF THE "SMOOTH" PARTS B^H , I^{α_1} , AND I^{α_2}** .
- The quadratic variation is independent of probabilistic properties.

TOY MODEL'S VOLATILITY

IMPLIED VOLATILITY

- The options' prices with toy-model are given by **REPLICATION PRICES**.
- So, the **IMPLIED VOLATILITY IS THE QUADRATIC VARIATION σ^2** .
- The **IMPLIED VOLATILITY IS INDEPENDENT OF THE "SMOOTH" PARTS B^H , I^{α_1} , AND I^{α_2}** .
- The quadratic variation is independent of probabilistic properties.
- Probability is irrelevant in option pricing and replication.

TOY MODEL'S VOLATILITY

ESTIMATING VOLATILITY

- So, implied volatility is quadratic variation.

TOY MODEL'S VOLATILITY

ESTIMATING VOLATILITY

- So, implied volatility is quadratic variation.
- A naïve approach to estimate the implied volatility historically would then be to use

$$\hat{\sigma}_n^2 = \frac{1}{2^n} \sum_{k=1}^{2^n} \left(R_{kT/2^n} - \frac{1}{2^n} \sum_{k=1}^{2^n} R_{kT/2^n} \right)^2 .$$

TOY MODEL'S VOLATILITY

ESTIMATING VOLATILITY

- So, implied volatility is quadratic variation.
- A naïve approach to estimate the implied volatility historically would then be to use

$$\hat{\sigma}_n^2 = \frac{1}{2^n} \sum_{k=1}^{2^n} \left(R_{kT/2^n} - \frac{1}{2^n} \sum_{k=1}^{2^n} R_{kT/2^n} \right)^2.$$

- So, in estimating variance one lets time go to infinity, while in estimating quadratic variation one lets time-increments go to zero.

TOY MODEL'S VOLATILITY

ESTIMATING VOLATILITY

- So, implied volatility is quadratic variation.
- A naïve approach to estimate the implied volatility historically would then be to use

$$\hat{\sigma}_n^2 = \frac{1}{2^n} \sum_{k=1}^{2^n} \left(R_{kT/2^n} - \frac{1}{2^n} \sum_{k=1}^{2^n} R_{kT/2^n} \right)^2.$$

- So, in estimating variance one lets time go to infinity, while in estimating quadratic variation one lets time-increments go to zero.
- With financial time series the naïve historical volatility estimation does not work: There is no price process in the microscopic level!

OUTLINE

1 TOY MODEL

2 HEDGING WITH QUADRATIC VARIATION

3 TOY MODEL'S VOLATILITY

4 REFERENCES

REFERENCES

[Cont](#) (2001): Empirical properties of asset returns: stylized facts and statistical issues.

[Föllmer](#) (1981): Calcul d'Itô sans probabilités.

[Schoenmakers, Kloeden](#) (1999): Robust Option Replication for a Black–Scholes Model Extended with Nondeterministic Trends.

[Russo, Vallois](#) (1993): Forward, backward and symmetric stochastic integration.

[Sottinen, Valkeila](#) (2003): On arbitrage and replication in the Fractional Black–Scholes pricing model.

This talk: [Bender, Sottinen, Valkeila](#) (2008): Pricing by hedging and no-arbitrage beyond semimartingales.