

4. Examples

Example 4.1 Implementation of the Example 3.1 in SAS. In SAS we can use the Proc Model procedure.

Simulate data from t -distribution with $\nu = 6$.
SAS:

```
data tdist;
do i = 1 to 500;
  y = tinv(ranuni(1258),6);
  z = 1;
  output;
end;
run;

proc model data = tdist;
  endogenous y;
  parms nu 5;
  eq.h1 = y**2 - nu/(nu-2);
  eq.h2 = y**4 - 3*nu**2/((nu-2)*(nu-4));
  fit h1 h2/gmm;
  instruments z / noint;
run;
```

Results:

The MODEL Procedure
Nonlinear GMM Summary of Residual Errors

Equation	DF	DF	SSE	MSE	Root MSE
	Model	Error			
h1	0.5	499.5	5671.5	11.3544	3.3696
h2	0.5	499.5	8834480	17686.6	133.0

Nonlinear GMM Parameter Estimates

Parameter	Estimate	Approx Std Err	t Value	Approx Pr > t
nu	5.374285	0.3931	13.67	<.0001

Number of Observations		Statistics for System	
Used	500	Objective	0.004917
Missing	0	Objective*N	2.4584

Thus $\hat{\nu} = 5.37$ with standard error 0.339. $J = 2.4584$ with 1 degree of freedom and p -value 0.117. Thus the data does not reject the moment conditions implied by the model.

Example 4.2 Normality of SP500 (daily) returns, y_t .
If normality holds, the moment conditions:

$$\begin{aligned}y_t - \mu &= 0 \\(y_t - \mu)^2 - \sigma^2 &= 0 \\(y_t - \mu)^3 / \sigma^3 &= 0 \\(y_t - \mu)^4 / \sigma^4 - 3 &= 0\end{aligned}$$

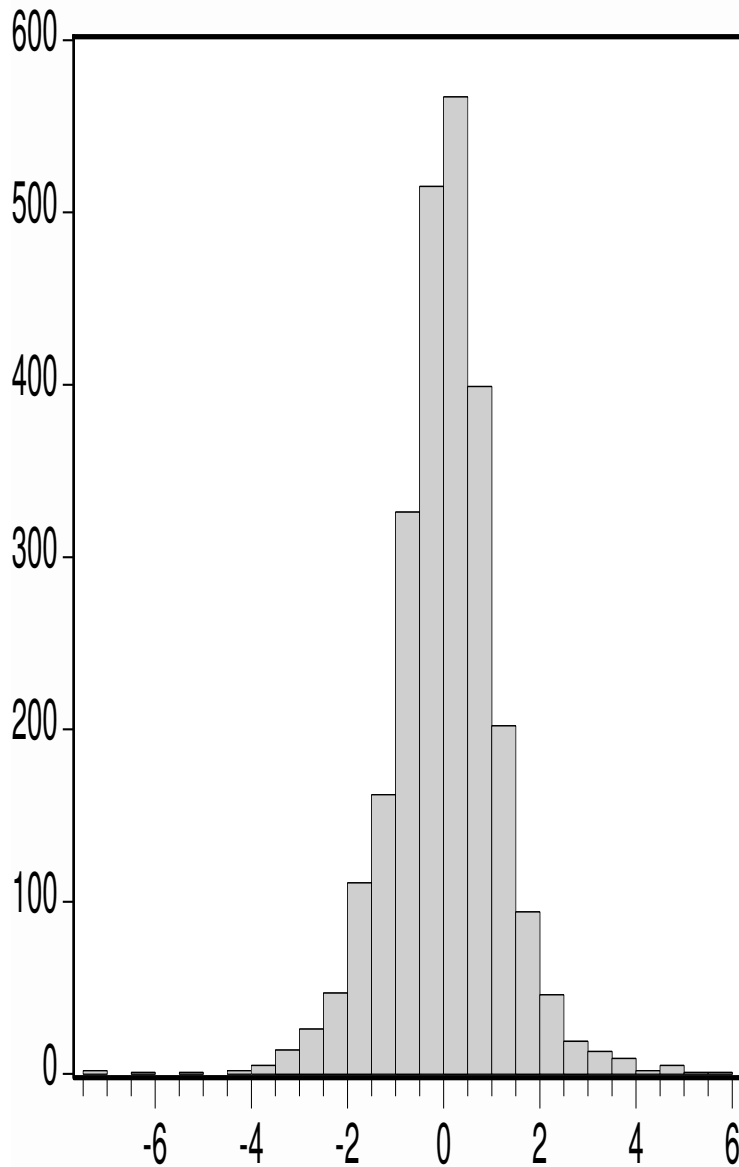
where $\mu = E[y_t]$, $\sigma^2 = \text{Var}[y_t]$. Data is obtained from finance.yahoo.com web site with sample period Jan 2, 1995 to May 19, 2005.

In EViews open Object \rightarrow New object ... \rightarrow System \rightarrow System

and write commands (c(1) = mean, c(2) = standard deviation)

```
@inst c
param c(1) 0 c(2) 1.0
sprete - c(1)
(sprete - c(1))^2 - c(2)^2
((sprete - c(1))/c(2))^3
((sprete - c(1))/c(2))^4 - 3
```

Sample statistics



Series: SPRET
Sample 2/01/1995 4/19/2005
Observations 2570

Mean	0.034876
Median	0.052299
Maximum	5.573247
Minimum	-7.113885
Std. Dev.	1.140887
Skewness	-0.110042
Kurtosis	6.117582

Jarque-Bera	1045.963
Probability	0.000000

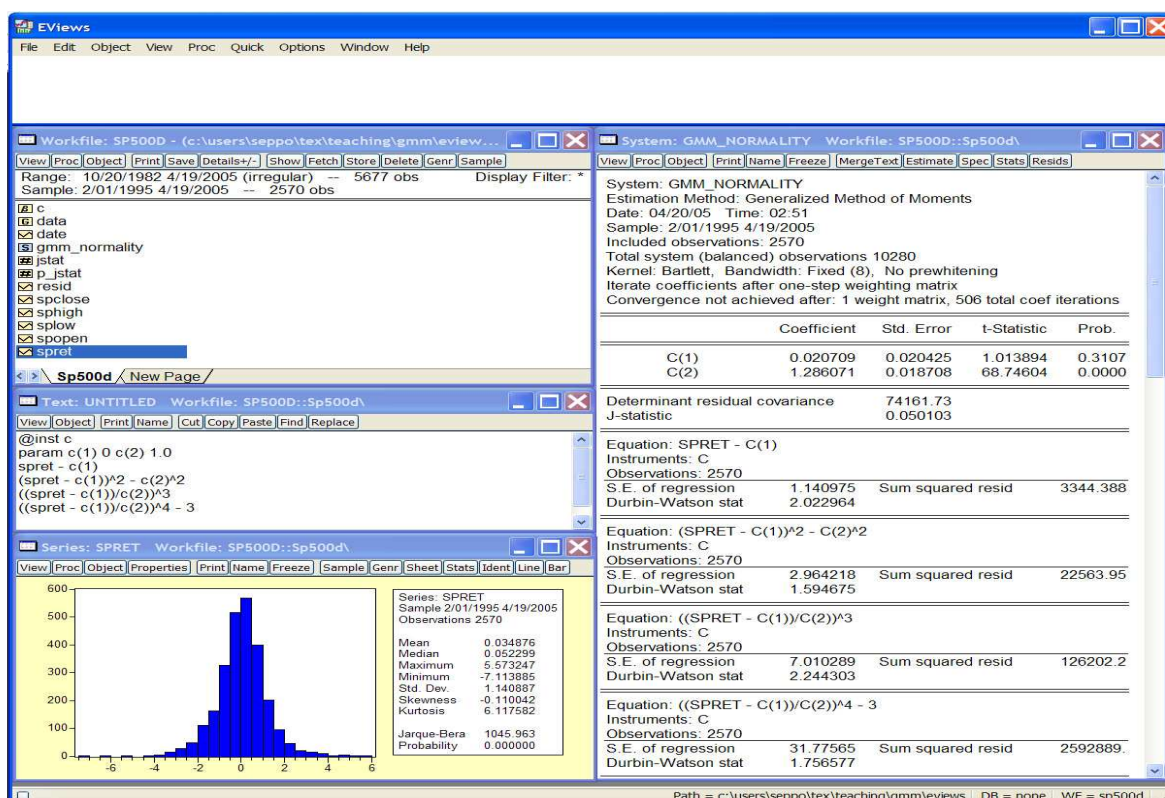
Select Estimate → GMM to get results

System: GMM_NORMALITY Estimation Method: Generalized Method of Moments Date: 04/20/05 Time: 02:51 Sample: 2/01/1995 4/19/2005 Included observations: 2570 Total system (balanced) observations 10280 Kernel: Bartlett, Bandwidth: Fixed (8), No prewhitening Iterate coefficients after one-step weighting matrix Convergence not achieved after: 1 weight matrix, 506 total coef iterations				
	Coefficient	Std. Error	t-Statistic	Prob.
C(1)	0.020709	0.020425	1.013894	0.3107
C(2)	1.286071	0.018708	68.74604	0.0000
Determinant residual covariance		74161.73		
J-statistic		0.050103		
Equation: SPRET - C(1)				
Instruments: C				
Observations: 2570				
S.E. of regression	1.140975	Sum squared resid	3344.388	
Durbin-Watson stat	2.022964			
Equation: (SPRET - C(1))^2 - C(2)^2				
Instruments: C				
Observations: 2570				
S.E. of regression	2.964218	Sum squared resid	22563.95	
Durbin-Watson stat	1.594675			
Equation: ((SPRET - C(1))/C(2))^3				
Instruments: C				
Observations: 2570				
S.E. of regression	7.010289	Sum squared resid	126202.2	
Durbin-Watson stat	2.244303			
Equation: ((SPRET - C(1))/C(2))^4 - 3				
Instruments: C				
Observations: 2570				
S.E. of regression	31.77565	Sum squared resid	2592889.	
Durbin-Watson stat	1.756577			

Note that the J -statistic in EViews is not multiplied by number of observations. Thus

$$J = 2570 \times 0.050103 \approx 128.8,$$

which is highly statistically significant ($df = 4 - 2 = 2$), and thus rejects the normality hypothesis.



Let us next test whether a t -distribution with location parameter μ , scale parameter σ^2 , and degrees of freedom parameter ν fits better. The density function is

$$f(y) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\pi\nu\sigma^2} \Gamma\left(\frac{\nu}{2}\right)} \left(1 + \frac{(y-\mu)^2}{\sigma^2(\nu-2)}\right)^{-\frac{1}{2}(\nu+1)}$$

with

$$E[y] = \mu,$$

$$\text{Var}[y] = \frac{\nu\sigma^2}{\nu-2},$$

and

$$E[(y-\mu)^4] = \frac{3\nu^2\sigma^4}{(\nu-2)(\nu-4)}.$$

The implied moment conditions are

$$E[y - \mu] = 0$$

$$E\left[(y - \mu)^2 - \sigma^2 \frac{\nu}{\nu-2}\right] = 0$$

$$E[(y - \mu)^3] = 0$$

$$E\left[(y - \mu)^4 - \frac{3\nu^2\sigma^4}{(\nu-2)(\nu-4)}\right] = 0$$

EViews estimation produces with commands (c(1) = mean, c(2) = scale, c(3) = df)

```
@inst c
param c(1) 0 c(2) 1.0 c(3) 7.0
sprete - c(1)
(sprete - c(1))^2 - c(2)^2*c(3)/(c(3)-2)
(sprete - c(1))^3
(sprete - c(1))^4 - (3*c(2)^4)*(c(3)^2)/((c(3)-2)*(c(3)-4))
```

System: GMM_T				
Estimation Method: Generalized Method of Moments				
Date: 04/25/05 Time: 00:29				
Sample: 2/01/1995 4/19/2005				
Included observations: 2570				
Total system (balanced) observations 10280				
Kernel: Bartlett, Bandwidth: Fixed (8), No prewhitening				
Iterate coefficients after one-step weighting matrix				
Convergence achieved after: 1 weight matrix, 8 total coef iterations				
	Coefficient	Std. Error	t-Statistic	Prob.
C(1)	0.040571	0.019412	2.089975	0.0366
C(2)	0.933509	0.029089	32.09109	0.0000
C(3)	6.132127	0.436527	14.04754	0.0000
Determinant residual covariance		2346223.		
J-statistic		0.000233		
Equation: SPRET - C(1)				
Instruments: C				
Observations: 2570				
S.E. of regression	1.140902	Sum squared resid	3343.955	
Durbin-Watson stat	2.023225			
Equation: (SPRET - C(1))^2 - C(2)^2*C(3)/(C(3)-2)				
Instruments: C				
Observations: 2570				
S.E. of regression	2.945789	Sum squared resid	22275.58	
Durbin-Watson stat	1.616382			
Equation: (SPRET - C(1))^3				
Instruments: C				
Observations: 2570				
S.E. of regression	14.94928	Sum squared resid	574122.4	
Durbin-Watson stat	2.241108			
Equation: (SPRET - C(1))^4 - (3*C(2)^4)*(C(3)^2)/((C(3)-2)*(C(3)-4))				
Instruments: C				
Observations: 2570				
S.E. of regression	87.50198	Sum squared resid	19654482	
Durbin-Watson stat	1.761688			

The J -statistic is $J = T \times J_{EViews} = 2570 \times 0.000233 = 0.598$ with p -value 0.439 ($df = 1$), which indicates

the return distribution seems to behave like a the t -distribution at least up to the first four moments with estimates $\hat{\mu} = 0.041$, $\hat{\sigma} = 0.934$, and $\hat{\nu} = 6.132$.

Example 4.4 Estimation of Dynamic Rational Expectations Model.

Hansen and Singleton (1982). Generalized instrumental variables method of nonlinear rational expectations models. *Econometrica* 50, 1269–1286. Errata: *Econometrica* 52, 267–268.

Theoretical background, data and SAS-code for a similar problem can be found from.

<http://support.sas.com/rnd/app/examples/ets/harvey/index.htm>